# Math 461 Spring 2024 

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## Outline

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2 4.7 Poisson random variables

3 4.8 Other Discrete Probability Distributions

HW4 is due Friday, 02/16, before the end of class. Please submit your HW4 as ONE pdf file via the HW4 folder in the course Moodle page.

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Solutions to HW3 is on my homepage.

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## (1) General Info

## 2 4.7 Poisson random variables

3 4.8 Other Discrete Probability Distributions

If $n$ independent trials, each results in a success with probability $p$, are performed, when $n$ is big, $p$ is small so that $n p$ is of moderate size, the number of successes in the $n$ trials is approximately a Poisson random variable with parameter $\lambda=n p$.

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## Examples

1 Number of misprints on a page of a book.
2 Number of people in a community over the age of 95 .

## Example

A machine produces screws, $1 \%$ of which are defective. Find the probability that in a box of 100 screws there are at most 3 defective ones. Assume independence.

## The number of defectives in the box is a binomial random variable

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## Example

A machine produces screws, $1 \%$ of which are defective. Find the probability that in a box of 100 screws there are at most 3 defective ones. Assume independence.

The number of defectives in the box is a binomial random variable with parameters $(100,0.01)$. So the exact answer is

$$
\begin{aligned}
P(X \leq 3)= & (0.99)^{100}+100 \cdot(0.01)(0.99)^{99} \\
& +\binom{100}{2}(0.01)^{2}(0.99)^{98}+\binom{100}{3}(0.01)^{3}(0.99)^{97} .
\end{aligned}
$$

$X$ is approximately a Poisson random variable with parameter 1 , so

$$
P(X \leq 3) \approx e^{-1}+e^{-1}+e^{-1} \frac{1}{2}+e^{-1} \frac{1}{6} .
$$

Poisson random variables also arise in situations where "incidents" occur at certain points in time, like earthquakes, people entering a certain establishment.


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In a lot of situations, the following assumptions are (approximately) satisfied: For some $\lambda>0$, the following hold:

1 The probability of exactly 1 incident occurs in a given interval of length $h$ is $\lambda h+o(h)$,
2 The probability of 2 or more incidents occur in an interval of length $h$ is $o(h)$.
3 For any integer $n \geq 1$, any non-negative integers $j_{1}, \ldots, j_{n}$, and any set of $n$ non-overlapping intervals, if $E_{i}$ denotes the event that exactly $j_{i}$ incidents occur in the $i$-th interval, $i=1, \ldots, n$, then $E_{1}, \ldots, E_{n}$ are independent.

Under the assumptions above, the number of incidents occurring in any interval of length $t$ is a Poisson random variable with parameter $\lambda t$. It suffices to deal with the case when the interval is $[0, t]$.


The event $\{N(t)=k\}$ can be written as the disjoint union of 2 events
$\square$

Under the assumptions above, the number of incidents occurring in any interval of length $t$ is a Poisson random variable with parameter $\lambda t$. It suffices to deal with the case when the interval is $[0, t]$.

Let $N(t)$ be the number of incidents occurring in $[0, t]$. For any $n \geq 1$, we divide $[0, t]$ into $n$ sub-intervals of equal length:


The event $\{N(t)=k\}$ can be written as the disjoint union of 2 events $A$ and $B$ where
$A$ is the event that " $k$ of the $n$ sub-intervals contains exactly 1 incident each and the other $n-k$ sub-intervals contains 0 incident", and $B$ is the event that " $N(t)=k$ and at least one of the sub-intervals contain 2 or more incidents". Thus $P(N(t)=k)=P(A)+P(B)$.
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$$
\begin{aligned}
P(B) & \leq P(\text { at least } 1 \text { subinterval contain } 2 \text { or more incidents }) \\
& =P\left(\cup_{i=1}^{n}\{\text { the } i \text {-th subinterval contain } 2 \text { or more incidents })\right. \\
& \leq \sum_{i=1}^{n} P(\text { the } i \text {-th subinterval contain } 2 \text { or more incidents }) \\
& =\sum_{i=1}^{n} o\left(\frac{t}{n}\right)=n \cdot o\left(\frac{t}{n}\right) \rightarrow 0
\end{aligned}
$$

as $n \rightarrow \infty$.

$$
P(A)=\binom{n}{k}\left(\frac{\lambda t}{n}+o\left(\frac{\lambda t}{n}\right)\right)^{k}\left(1-\frac{\lambda t}{n}-o\left(\frac{\lambda t}{n}\right)\right)^{n-k} .
$$

Since

$$
n\left(\frac{\lambda t}{n}+o\left(\frac{\lambda t}{n}\right)\right) \rightarrow \lambda t
$$

we have

$$
P(A) \rightarrow e^{-\lambda t} \frac{(\lambda t)^{k}}{k!}
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Thus

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## Examples

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## Examples

(a) The number of earthquakes during some fixed time interval.
(b) The number of $\alpha$-particles discharged from some radioactive material in a fixed period of time.

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## (1) General Info

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Suppose that independent trials, each results in a success with probability $p \in(0,1)$ and a failure with probability $1-p$, are performed until a success occurs. Let $X$ be the number of trials need, then

$$
P(X=n)=(1-p)^{n-1} p, \quad n=1,2, \ldots
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Such a random variable is called a geometric random variable with parameter $p$.

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Such a random variable is called a geometric random variable with parameter $p$.

If $X$ is a geometric random variable with parameter $p$, then for any $k \geq 1$,

$$
P(X \geq k)=(1-p)^{k-1} .
$$

## Example 1

Cards are randomly selected from an ordinary deck, one at a time, until a spade is obtained. If we assume that each card is returned to the deck before the next one is selected, find the probability that (a) exactly 10 cards are needed; (b) at least 10 cards are needed.

## Example 1

Cards are randomly selected from an ordinary deck, one at a time, until a spade is obtained. If we assume that each card is returned to the deck before the next one is selected, find the probability that (a) exactly 10 cards are needed; (b) at least 10 cards are needed.

## Solution

The number of cards needed is a geometric random variable with parameter $\frac{1}{4}$. Thus (a) $\left(\frac{3}{4}\right)^{9}\left(\frac{1}{4}\right)$; (b) $\left(\frac{3}{4}\right)^{9}$.
$X$ is a geometric random variable with parameter $p$, then

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Let $q=1-p$. Then

$$
\begin{aligned}
E[X] & =\sum_{n=1}^{\infty} n q^{n-1} p=p \sum_{n=0}^{\infty} \frac{d}{d q}\left(q^{n}\right) \\
& =p \frac{d}{d q}\left(\sum_{n=0}^{\infty} q^{n}\right)=p \frac{d}{d q}\left(\frac{1}{1-q}\right) \\
& =\frac{p}{(1-q)^{2}}=\frac{1}{p} .
\end{aligned}
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E\left[X^{2}\right] & =\sum_{n=1}^{\infty} n^{2} q^{n-1} p=p \sum_{n=0}^{\infty} \frac{d}{d q}\left(n q^{n}\right) \\
& =p \frac{d}{d q}\left(\sum_{n=0}^{\infty} n q^{n}\right)=p \frac{d}{d q}\left(\frac{q}{1-q} E[X]\right) \\
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Thus

$$
\operatorname{Var}(X)=E\left[X^{2}\right]-(E[X])^{2}=\frac{1-p}{p^{2}}
$$

