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Math 461 Spring 2024

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University of Illinois Urbana-Champaign

January 29, 2024

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Outline

3.2 Conditional Probabilities

3.3 Bayes' Formula

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2 3.2 Conditional Probabilities

3.3 Bayes' Formula

Solutions to HW1 are available on my hmepage.

HW2 is due this Friday at the end of the class.

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If P(F) > 0, we define

$$P(E|F) = rac{P(E \cap F)}{P(F)}.$$

If P(F) = 0, P(E|F) is undefined.

Using the definition of conditional probability, one can easily check $P(E \cap F) = P(F)P(E|F).$

More generally, we have

 $P(\bigcap_{i=1}^{n} E_i) = P(E_1)P(E_2|E_1)\cdots P(E_n|\bigcap_{i=1}^{n-1} E_i).$

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Example 3

Suppose that a box contains 8 red balls and 4 white balls. We randomly draw two balls from the box without replacement. Find the probability that (a) both balls are red; (b) the second ball is red.

(a) Using the obvious notation, $P(R_1 \cap R_2) = P(R_1)P(R_2|R_1) = \frac{8}{12} \frac{7}{11}.$

$$P(R_2) = P(R_1 \cap R_2) + P(W_1 \cap R_2) = \frac{8}{12} \frac{7}{11} + \frac{4}{12} \frac{8}{11}.$$

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Example 3

Suppose that in the previous example. 3 balls are randomly selected from the box without replacement. Find the probability that all 3 are red.

$P(R_1 \cap R_2 \cap R_3) = P(R_1)P(R_2|R_1)P(R_3|R_1 \cap R_2)$ $= \frac{8}{12}\frac{7}{11}\frac{6}{10}.$

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Example 4

An ordinary deck of 52 cards is randomly divided into 4 distinct piles of 13 each. Find the probability that each pile has exactly 1 ace.

Solution. From an example in Section 2.5, we know that the answer is

$$\frac{4!\binom{48}{12,12,12,12}}{\binom{52}{13,13,13}}$$

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Example 4

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Solution. From an example in Section 2.5, we know that the answer is

$$\frac{4!\binom{48}{12,12,12,12}}{\binom{52}{13,13,13,13}}$$

Solution by conditional probability. For i = 1, 2, 3, 4, let E_i be the event that the *i*-th pile has exactly 1 ace. Then

. (19)

$$P(E_1) = \frac{4\binom{12}{12}}{\binom{52}{13}}, \quad P(E_2|E_1) = \frac{3\binom{12}{12}}{\binom{39}{13}},$$
$$P(E_3|E_1 \cap E_2) = \frac{2\binom{24}{12}}{\binom{26}{13}}, \quad P(E_4|E_1 \cap E_2 \cap E_3) = \frac{\binom{12}{12}}{\binom{13}{13}} = 1.$$

- (26)

So the answer is

$4\binom{48}{12}$	$3\binom{36}{12}$	$2\binom{24}{12}$	$\binom{12}{12}$
⁽⁵²)	$(39 \\ 13)$	(²⁶)	$\overline{\begin{pmatrix}13\\13\end{pmatrix}}$.

The 2 answers are the same. See the book for yet another solution via conditional probability.

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$$P(E_1) = \frac{4\binom{48}{12}}{\binom{52}{13}}, \quad P(E_2|E_1) = \frac{3\binom{36}{12}}{\binom{39}{13}},$$
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So the answer is
$$4\binom{48}{12} 3\binom{36}{12} 2\binom{24}{12} \binom{12}{12}$$

 $\begin{array}{c|c}\hline \hline \begin{pmatrix} 39\\13 \end{pmatrix} \hline \begin{pmatrix} 26\\13 \end{pmatrix} \hline \begin{pmatrix} 13\\13 \end{pmatrix} \\ \hline \end{pmatrix} .$

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 $\binom{52}{13}$

3.2 Conditional Probabilities

3.3 Bayes' Formula ●000000000

Outline



2 3.2 Conditional Probabilities



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Example 1

A certain blood test is 95% effective in detecting a certain disease when it is in fact present. However, the test also yields a "false positive" result for 1% of the healthy people tested. If 0.5% of the population has the disease, what is the probability that a person has the disease given that the person's test result is positive?

Solution. Let *E* be the event that the person has the disease, and *F* the event that the person's test result is positive. We are looking for P(E|F), which is equal to

$$\frac{P(E\cap F)}{P(F)}.$$

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3.2 Conditional Probabilities

3.3 Bayes' Formula 00●00000000

We are given $P(E) = 0.005, P(E^c) = 0.995$ and $P(F|E) = .95 \quad P(F|E^c) = .01.$ Thus $P(E \cap F) = P(E)P(F|E) = (0.005) \cdot (0.95)$ and $P(F) = P(E \cap F) + P(E^c \cap F) = P(E)P(F|E) + P(E^c)P(F|E^c)$ $= (0.005) \cdot (0.95) + (0.995) \cdot (0.01).$ The answer is $(0.005) \cdot (0.95)$ $(0.005) \cdot (0.95) + (0.995) \cdot (0.01) \approx 0.323.$

The example above is a special case of the following general situation. Suppose $A_1, A_2, ..., A_n$ are *n* disjoint events with their union being the whole sample space and with $P(A_i) > 0$ for each i = 1, ..., n. Let *B* be an event with P(B) > 0. Suppose that $P(A_i), P(B|A_i), i = 1, ..., n$ are given. Find $P(A_i|B)$.

$$B = B \cap \left(\bigcup_{j=1}^{n} A_j \right) = \bigcup_{j=1}^{n} (B \cap A_j).$$
$$P(B) = \sum_{j=1}^{n} P(A_j) P(B|A_j).$$

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Thus

$$P(A_i|B) = rac{P(A_i \cap B)}{P(B)} = rac{P(A_i)P(B|A_i)}{\sum_{j=1}^n P(A_j)P(B|A_j)}.$$

The formula above is known as the Bayes' formula. You do not need to memorize this formula. It is much easier to remember the short derivation of it.

Example 2

In answering a certain multiple choice question with 5 possible answers, a student either knows the answer or guesses. Assume that a student knows the answer with probability 0.8. Assume that, when not knowing the answer, the student guesses the 5 answers with equal probability. Find the probability that the student knows the answer given that the student answered it correctly.

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Solution. Let K be the event that the student knows the answer, and C the event that the student answered it correctly. Then

$$P(K) = 0.8, P(K^c) = 0.2$$

and

$$P(C|K) = 1$$
 $P(C|K^c) = 0.2$.

So

$$P(K|C) = \frac{P(K \cap C)}{P(C)} = \frac{P(K)P(C|K)}{P(K)P(C|K) + P(K^c)P(C|K^c)}$$
$$= \frac{(0.8) \cdot 1}{(0.8) \cdot 1 + (0.2) \cdot (0.2)}.$$

Example 3

Suppose that there are 3 chests of drawers and each chest has 2 drawers. The first chest has a gold coin in each drawer; the second chest has a gold in one drawer and a silver coin in the other; the third chest has a silver coin in each drawer. A chest is chosen at random and a drawer is randomly opened. If the drawer has a gold coin, what is the probability that the other drawer also has a glod coin?

Solution. For i = 1, 2, 3, let E_i be the event that the *i*-th chest is chosen, and let *G* be the event that the drawer opened has a gold coin. We are looking for $P(E_1|G)$.

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$$P(E_1) = P(E_2) = P(E_3) = \frac{1}{3};$$

and

$$P(G|E_1) = 1, \quad P(G|E_2) = \frac{1}{2}, \quad P(G|E_3) = 0.$$

So

$$P(E_1|G) = \frac{P(E_1 \cap G)}{P(E_1 \cap G) + P(E_2 \cap G) + P(E_3 \cap G)}$$

= $\frac{P(E_1)P(G|E_1)}{P(E_1)P(G|E_1) + P(E_2)P(G|E_2) + P(E_3)P(G|E_3)}$
= $\frac{\frac{1}{3}}{\frac{1}{3} + \frac{1}{3}\frac{1}{2}} = \frac{2}{3}.$

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A plane is missing, and it is presumed that it is equally likely to have gone down in any of 3 possible regions. Let $1 - \beta_i$ be the probability that the plane will be found upon a search of the region when the plane is, in fact, in that region, i = 2, 3. Find the probability that the plane is in the *i*-th region given that a search of region 1 did not locate the plane.

Solution. For i = 1, 2, 3, let E_i be the event that the plane is the *i*-th region. Let *F* be the event that a search of region 1 did not locate the plane. We are looking for $P(E_1|F)$, $P(E_2|F)$ and $P(E_3|F)$.

$$P(E_1) = P(E_2) = P(E_3) = \frac{1}{3},$$

and

 $P(F|E_1) = \beta_1, \quad P(F|E_2) = P(F|E_3) = 1.$

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Similarly,

$$P(E_2|F) = P(E_3|F) = \frac{1}{\beta_1 + 2}.$$

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