# Math 461 Spring 2024 

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## Outline

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## 2 3.2 Conditional Probabilities

3 3.3 Bayes' Formula

Solutions to HW1 are available on my hmepage.

## HW2 is due this Friday at the end of the class.

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## (1) General Info

## 2) 3.2 Conditional Probabilities

3 3.3 Bayes' Formula

## Definition

If $P(F)>0$, we define

$$
P(E \mid F)=\frac{P(E \cap F)}{P(F)} .
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If $P(F)=0, P(E \mid F)$ is undefined.

Using the definition of conditional probability, one can easily check

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More generally, we have

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P\left(\cap_{i=1}^{n} E_{i}\right)=P\left(E_{1}\right) P\left(E_{2} \mid E_{1}\right) \cdots P\left(E_{n} \mid \cap_{i=1}^{n-1} E_{i}\right) .
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These formulas are very useful in finding the probability of intersections.

## Example 3

Suppose that a box contains 8 red balls and 4 white balls. We randomly draw two balls from the box without replacement. Find the probability that (a) both balls are red; (b) the second ball is red.
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$$

$$
P\left(R_{2}\right)=P\left(R_{1} \cap R_{2}\right)+P\left(W_{1} \cap R_{2}\right)=\frac{8}{12} \frac{7}{11}+\frac{4}{12} \frac{8}{11} .
$$

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\begin{aligned}
P\left(R_{1} \cap R_{2} \cap R_{3}\right) & =P\left(R_{1}\right) P\left(R_{2} \mid R_{1}\right) P\left(R_{3} \mid R_{1} \cap R_{2}\right) \\
& =\frac{8}{12} \frac{7}{11} \frac{6}{10} .
\end{aligned}
$$

## Example 4

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Solution. From an example in Section 2.5

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An ordinary deck of 52 cards is randomly divided into 4 distinct piles of 13 each. Find the probability that each pile has exactly 1 ace.

Solution. From an example in Section 2.5, we know that the answer is

$$
\frac{4!\binom{48}{12,12,12,12}}{\binom{52}{13,13,13,13}}
$$

Solution by conditional probability. For $i=1,2,3,4$, let $E_{i}$ be the event that the $i$-th pile has exactly 1 ace. Then

$$
\begin{gathered}
P\left(E_{1}\right)=\frac{4\binom{48}{12}}{\binom{52}{13}}, \quad P\left(E_{2} \mid E_{1}\right)=\frac{3\binom{36}{12}}{\binom{39}{13}}, \\
P\left(E_{3} \mid E_{1} \cap E_{2}\right)=\frac{2\binom{24}{12}}{\binom{26}{13}}, \quad P\left(E_{4} \mid E_{1} \cap E_{2} \cap E_{3}\right)=\frac{\binom{12}{12}}{\binom{13}{13}}=1 .
\end{gathered}
$$

So the answer is

$$
\frac{4\binom{48}{12}}{\binom{52}{13}} \frac{3\binom{36}{12}}{\binom{39}{13}} \frac{2\binom{24}{12}}{\binom{12}{13}} \frac{\binom{13}{13}}{13} \text {. }
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P\left(E_{1}\right)=\frac{4\binom{48}{12}}{\binom{52}{13}}, \quad P\left(E_{2} \mid E_{1}\right)=\frac{3\binom{36}{12}}{\binom{39}{13}}, \\
P\left(E_{3} \mid E_{1} \cap E_{2}\right)=\frac{2\binom{24}{12}}{\binom{26}{13}}, \quad P\left(E_{4} \mid E_{1} \cap E_{2} \cap E_{3}\right)=\frac{\binom{12}{12}}{\binom{13}{13}}=1 .
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$$

The 2 answers are the same. See the book for yet another solution via conditional probability.

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## (1) General Info

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## Example 1

A certain blood test is $95 \%$ effective in detecting a certain disease when it is in fact present. However, the test also yields a "false positive" result for $1 \%$ of the healthy people tested. If $0.5 \%$ of the population has the disease, what is the probability that a person has the disease given that the person's test result is positive?


## Example 1

A certain blood test is $95 \%$ effective in detecting a certain disease when it is in fact present. However, the test also yields a "false positive" result for $1 \%$ of the healthy people tested. If $0.5 \%$ of the population has the disease, what is the probability that a person has the disease given that the person's test result is positive?

Solution. Let $E$ be the event that the person has the disease, and $F$ the event that the person's test result is positive. We are looking for $P(E \mid F)$, which is equal to

$$
\frac{P(E \cap F)}{P(F)} .
$$

We are given

$$
P(E)=0.005, \quad P\left(E^{c}\right)=0.995
$$

and

$$
P(F \mid E)=.95 \quad P\left(F \mid E^{c}\right)=.01
$$

Thus

$$
P(E \cap F)=P(E) P(F \mid E)=(0.005) \cdot(0.95)
$$

and

$$
\begin{aligned}
P(F) & =P(E \cap F)+P\left(E^{c} \cap F\right)=P(E) P(F \mid E)+P\left(E^{c}\right) P\left(F \mid E^{c}\right) \\
& =(0.005) \cdot(0.95)+(0.995) \cdot(0.01) .
\end{aligned}
$$

The answer is

$$
\frac{(0.005) \cdot(0.95)}{(0.005) \cdot(0.95)+(0.995) \cdot(0.01)} \approx 0.323 .
$$

The example above is a special case of the following general situation. Suppose $A_{1}, A_{2}, \ldots, A_{n}$ are $n$ disjoint events with their union being the whole sample space and with $P\left(A_{i}\right)>0$ for each $i=1, \ldots, n$. Let $B$ be an event with $P(B)>0$. Suppose that $P\left(A_{i}\right), P\left(B \mid A_{i}\right), i=1, \ldots, n$ are given. Find $P\left(A_{i} \mid B\right)$.

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$$
B=B \cap\left(\cup_{j=1}^{n} A_{j}\right)=\cup_{j=1}^{n}\left(B \cap A_{j}\right) .
$$

So

$$
P(B)=\sum_{j=1}^{n} P\left(A_{j}\right) P\left(B \mid A_{j}\right) .
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Thus

$$
P\left(A_{i} \mid B\right)=\frac{P\left(A_{i} \cap B\right)}{P(B)}=\frac{P\left(A_{i}\right) P\left(B \mid A_{i}\right)}{\sum_{j=1}^{n} P\left(A_{j}\right) P\left(B \mid A_{j}\right)} .
$$

The formula above is known as the Bayes' formula. You do not need to memorize this formula. It is much easier to remember the short derivation of it.


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## Example 2

In answering a certain multiple choice question with 5 possible answers, a student either knows the answer or guesses. Assume that a student knows the answer with probability 0.8 . Assume that, when not knowing the answer, the student guesses the 5 answers with equal probability. Find the probability that the student knows the answer given that the student answered it correctly.

Solution. Let $K$ be the event that the student knows the answer, and $C$ the event that the student answered it correctly. Then

$$
P(K)=0.8, \quad P\left(K^{c}\right)=0.2
$$

and

$$
P(C \mid K)=1 \quad P\left(C \mid K^{c}\right)=0.2
$$

So

$$
\begin{aligned}
P(K \mid C) & =\frac{P(K \cap C)}{P(C)}=\frac{P(K) P(C \mid K)}{P(K) P(C \mid K)+P\left(K^{c}\right) P\left(C \mid K^{c}\right)} \\
& =\frac{(0.8) \cdot 1}{(0.8) \cdot 1+(0.2) \cdot(0.2)}
\end{aligned}
$$

## Example 3

Suppose that there are 3 chests of drawers and each chest has 2 drawers. The first chest has a gold coin in each drawer; the second chest has a gold in one drawer and a silver coin in the other; the third chest has a silver coin in each drawer. A chest is chosen at random and a drawer is randomly opened. If the drawer has a gold coin, what is the probability that the other drawer also has a glod coin?

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Solution. For $i=1,2,3$, let $E_{i}$ be the event that the $i$-th chest is chosen, and let $G$ be the event that the drawer opened has a gold coin. We are looking for $P\left(E_{1} \mid G\right)$.

$$
P\left(E_{1}\right)=P\left(E_{2}\right)=P\left(E_{3}\right)=\frac{1}{3}
$$

and

$$
P\left(G \mid E_{1}\right)=1, \quad P\left(G \mid E_{2}\right)=\frac{1}{2}, \quad P\left(G \mid E_{3}\right)=0
$$

So

$$
\begin{aligned}
P\left(E_{1} \mid G\right) & =\frac{P\left(E_{1} \cap G\right)}{P\left(E_{1} \cap G\right)+P\left(E_{2} \cap G\right)+P\left(E_{3} \cap G\right)} \\
& =\frac{P\left(E_{1}\right) P\left(G \mid E_{1}\right)}{P\left(E_{1}\right) P\left(G \mid E_{1}\right)+P\left(E_{2}\right) P\left(G \mid E_{2}\right)+P\left(E_{3}\right) P\left(G \mid E_{3}\right)} \\
& =\frac{\frac{1}{3}}{\frac{1}{3}+\frac{1}{3} \frac{1}{2}}=\frac{2}{3} .
\end{aligned}
$$

A plane is missing, and it is presumed that it is equally likely to have gone down in any of 3 possible regions. Let $1-\beta_{i}$ be the probability that the plane will be found upon a search of the region when the plane is, in fact, in that region, $i=, 2,3$. Find the probability that the plane is in the $i$-th region given that a search of region 1 did not locate the plane.

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Solution. For $i=1,2,3$, let $E_{i}$ be the event that the plane is the $i$-th region. Let $F$ be the event that a search of region 1 did not locate the plane. We are looking for $P\left(E_{1} \mid F\right), P\left(E_{2} \mid F\right)$ and $P\left(E_{3} \mid F\right)$.

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$$
P\left(E_{1}\right)=P\left(E_{2}\right)=P\left(E_{3}\right)=\frac{1}{3}
$$

and

$$
P\left(F \mid E_{1}\right)=\beta_{1}, \quad P\left(F \mid E_{2}\right)=P\left(F \mid E_{3}\right)=1 .
$$

So

$$
\begin{aligned}
P\left(E_{1} \mid F\right) & =\frac{P\left(E_{1} \cap F\right)}{P\left(E_{1} \cap F\right)+P\left(E_{2} \cap F\right)+P\left(E_{3} \cap F\right)} \\
& =\frac{P\left(E_{1}\right) P\left(F \mid E_{1}\right)}{P\left(E_{1}\right) P\left(F \mid E_{1}\right)+P\left(E_{2}\right) P\left(F \mid E_{2}\right)+P\left(E_{3}\right) P\left(F \mid E_{3}\right)} \\
& =\frac{\beta_{1}}{\beta_{1}+2} .
\end{aligned}
$$

So

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\begin{aligned}
P\left(E_{1} \mid F\right) & =\frac{P\left(E_{1} \cap F\right)}{P\left(E_{1} \cap F\right)+P\left(E_{2} \cap F\right)+P\left(E_{3} \cap F\right)} \\
& =\frac{P\left(E_{1}\right) P\left(F \mid E_{1}\right)}{P\left(E_{1}\right) P\left(F \mid E_{1}\right)+P\left(E_{2}\right) P\left(F \mid E_{2}\right)+P\left(E_{3}\right) P\left(F \mid E_{3}\right)} \\
& =\frac{\beta_{1}}{\beta_{1}+2} .
\end{aligned}
$$

Similarly,

$$
P\left(E_{2} \mid F\right)=P\left(E_{3} \mid F\right)=\frac{1}{\beta_{1}+2} .
$$

